

Genome-wide association study in a Chinese Han population identifies nine new susceptibility loci for systemic lupus erythematosus

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We performed a genome-wide association study (GWAS) of systemic lupus erythematosus (SLE) in a Chinese Han population by genotyping 1,047 cases and 1,205 controls using Illumina Human610-Quad BeadChips and replicating 78 SNPs in two additional cohorts (3,152 cases and 7,050 controls). We identified nine new susceptibility loci (*ETS1*, *IKZF1*, *RASGRP3*, *SLC15A4*, *TNIP1*, 7q11.23, 10q11.22, 11q23.3 and 16p11.2; $1.77 \times 10^{-25} \leq P_{\text{combined}} \leq 2.77 \times 10^{-8}$) and confirmed seven previously reported loci (*BLK*, *IRF5*, *STAT4*, *TNFAIP3*, *TNFSF4*, 6q21 and 22q11.21; $5.17 \times 10^{-42} \leq P_{\text{combined}} \leq 5.18 \times 10^{-12}$). Comparison with previous GWAS findings highlighted the genetic heterogeneity of SLE susceptibility between Chinese Han and European populations. This study not only advances our understanding of the genetic basis of SLE but also highlights the value of performing GWAS in diverse ancestral populations.

Systemic lupus erythematosus (SLE) is a prototypic systemic autoimmune disease that is characterized by a diverse array of

autoantibody production, complement activation, immune complex deposition and tissue and organ damage¹ and that is influenced by both genetic and environmental factors². SLE affects predominantly women (prevalence ratio of women to men is 9:1), particularly during the childbearing years. There are marked disparities in SLE incidence and prevalence worldwide; SLE prevalence varies among different ethnic and geographical populations³. The prevalence of SLE ranges from 31 to 70 cases per 100,000 persons among Chinese populations⁴ and from 7 to 71 cases per 100,000 persons in European populations⁵, and lupus nephritis is more prevalent in Chinese populations than in European populations³. The ethnic and genetic heterogeneity of SLE may contribute to the complexity of its clinical manifestation.

Over the past two decades, numerous studies have identified multiple genetic factors related to SLE. In particular, four recent genome-wide association studies (GWAS) of SLE in European populations have identified more than 20 robust susceptibility genes and/or loci⁶⁻⁹. The genetic heterogeneity between ethnic populations has

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Table 1 Summary of samples used in GWAS and replication studies

| | Cases | | | | | Controls | | | |
|----------------------------|-------------|-----------------|--------------------------|-------------|------------|-------------|-----------------|-------------|------------|
| | Sample size | Mean age (s.d.) | Mean age of onset (s.d.) | Male/female | Female (%) | Sample size | Mean age (s.d.) | Male/female | Female (%) |
| GWAS ^a | 1,047 | 34.02 (11.53) | 29.81 (10.08) | 63/984 | 93.98 | 1,205 | 34.75 (12.97) | 673/532 | 44.15 |
| Replication 1 ^b | 1,643 | 35.36 (12.10) | 30.89 (11.21) | 136/1,507 | 91.72 | 5,930 | 29.64 (11.31) | 2,729/3,201 | 53.98 |
| Replication 2 ^c | 1,509 | 32.85 (11.16) | 28.18 (10.42) | 113/1,396 | 92.51 | 1,120 | 32.23 (14.48) | 415/705 | 62.95 |
| Total | 4,199 | 34.12 (11.68) | 29.65 (10.96) | 312/3,887 | 92.57 | 8,255 | 30.74 (12.18) | 3,817/4,438 | 53.76 |

^aGWAS samples are from central China. ^bReplication 1 samples are from central China. ^cReplication 2 sample are from southern China.

been suggested to be important in SLE risk¹⁰, showing the need for further GWAS in non-European populations.

We conducted a GWAS of SLE in a Chinese Han population by analyzing 493,955 autosomal SNPs in 1,047 affected individuals (cases) and 1,205 controls (Table 1). Principal component analysis and genomic control ($\lambda_{gc} = 1.045$ for all 493,955 SNPs and 1.039 after excluding 4,840 SNPs within the major histocompatibility complex (MHC) region) indicated minimal inflation of the GWA results due to population stratification (Supplementary Fig. 1).

The analysis revealed association at three loci—6p21 (MHC)^{6,7,9}, 2q32.3 (*STAT4*)^{6,9} and 8p23.1 (*BLK*)⁶—with genome-wide significance ($P < 5 \times 10^{-8}$) (Fig. 1a). Furthermore, a quantile-quantile plot of the observed P values showed a clear deviation at the tail of the distribution from the null distribution (the distribution expected if there were no association) even after 4,840 SNPs were removed from the MHC region. This suggests that the observed P values, particularly the ones within the tail of the distribution, are smaller than those expected by chance and probably reflect true genetic association (Fig. 1b).

In the MHC region, 13 SNPs, all located within the HLA class II region, showed genome-wide significant association ($P < 5 \times 10^{-8}$); the most significant association was identified at rs9271100 ($P = 1.42 \times 10^{-12}$, odds ratio (OR) = 1.9, 95% CI = 1.59–2.27) (Supplementary Fig. 2). Further conditional association analysis of the 13 SNPs confirmed that there were two independent associations within the HLA class II region at rs9271100 near *HLA-DRB1* and at rs3997854 near *HLA-DQA2*. Our association results within the MHC region are

largely consistent with findings (including association at *HLA-DRB1*) identified in previous GWAS of SLE in European populations⁶.

We performed a replication study by genotyping 78 non-MHC SNPs (from 67 loci) in two additional cohorts of Chinese Han individuals (replication 1: 1,643 cases and 5,930 controls; replication 2: 1,509 cases and 1,120 controls; Table 1 and Supplementary Table 1). In the replication study, the association analysis was performed in the two individual replication samples separately as well as in the combined GWAS and replication sample. Twenty-one SNPs within 16 loci were validated, with independent evidence supporting association ($P < 0.03$) for the two replication samples and highly significant evidence in the combined sample that surpassed genome-wide significance ($P_{\text{combined}} < 5 \times 10^{-8}$). The 16 confirmed susceptibility loci were located at 1q25.1, 2p22.3, 2q32.3, 5q33.1, 6q21, 6q23.3, 7p12.2, 7q11.23, 7q32.1, 8p23.1, 10q11.22, 11q23.3, 11q24.3, 12q24.32, 16p11.2 and 22q11.21 ($5.17 \times 10^{-42} \leq P_{\text{combined}} \leq 2.77 \times 10^{-8}$) (Table 2). The associations at these 16 loci were independent of the associations within the MHC region because the associations at these loci remained similar even after controlling for the genetic effect of rs9271100 and rs3997854 within the MHC region (Supplementary Table 2). Pairwise interaction analysis was also performed among the top SNPs from the 16 non-MHC loci and the 2 MHC loci (rs9271100 and rs3997854), but no interaction was identified ($P > 0.05$ after correction for multiple testing; data not shown).

The associations at 1q25.1, 2q32.3, 6q21, 6q23.3, 7q32.1, 8p23.1 and 22q11.21 were reported by previous GWAS in European populations,

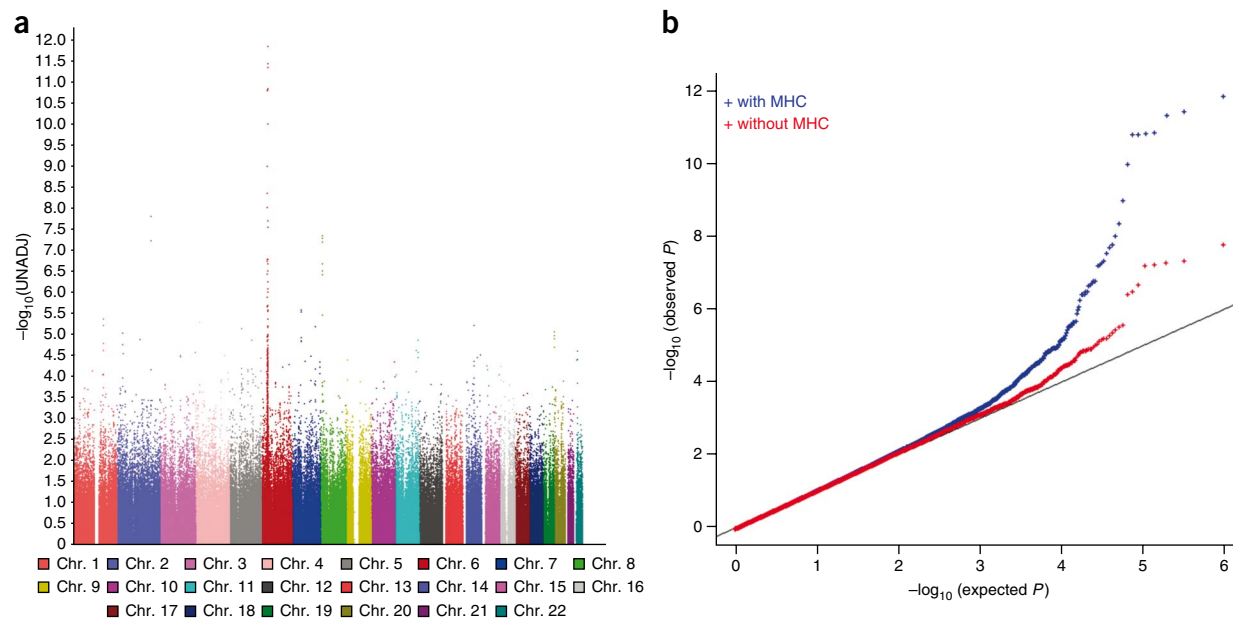


Figure 1 Summary of genome-wide association results for 1,047 cases and 1,205 controls. (a) The genome-wide P values ($-\log_{10} P$) of the logistic regression analysis adjusted for gender (493,955 SNPs) plotted against position on each chromosome. Each chromosome is depicted in a different color. (b) Quantile-quantile plots of the observed P values versus the expected values from P value of association. The plot in blue was based on the entire set of 493,955 SNPs, whereas the plot in red was obtained after removing 4,840 SNPs within MHC region (Chr. 6: 25–37 Mb).

associations at 10 SNPs within *BLK*, *IRF5*, *NMNAT2*, *STAT4*, *UBE2L3*, 1q25.1, 5q33.3 and 6q21 were also supported by our GWAS analysis with significant or suggestive evidence ($P < 0.01$) (**Supplementary Table 4a**, Group 1). However, our study did not provide evidence for the associations at the remaining 35 SNPs within *ATG5*, *BANK1*, *BLK*, *FCGR2A*, the MHC region, *ICA1*, *IRF5*, *ITGAM*, *PHRF1* (also known as *KIAA1542*), *LYN*, *PTPN22*, *PXK*, *SCUBE1*, *TNFAIP3* and *XKR6* ($P > 0.05$). Of these 35 SNPs, 24 SNPs (including 5 within the MHC region) are very rare (20 SNPs) or have low population frequency (4 SNPs) in the Chinese Han population (MAF $< 5\%$) (**Supplementary Table 4a**, Group 2 and 3). The rareness or low frequency of these 24 SNPs may be responsible for the failure to detect their associations in the Chinese Han population. Of note, the remaining 11 SNPs, including 2 SNPs within the HLA class II region identified by the International Consortium for Systemic Lupus Erythematosus Genetics (SLEGEN) study⁷, have a similar MAF in both the Chinese Han and European populations but failed to show any associations in the Chinese Han population (**Supplementary Table 4a**, Group 4). Conversely, none of the top 13 MHC SNPs identified in our GWAS analysis showed significant association in a GWAS of European populations⁶ (**Supplementary Table 4b**). Furthermore, of the 10 newly discovered non-MHC SNPs identified by this study, only rs10036748 within *TNIP1* showed significant association ($P = 10^{-4}$ – 10^{-3}) in European populations; this SNP also shows a much higher frequency (77%) in European populations than in Chinese populations (26%) (**Supplementary Table 4b**). None of the ten SNPs is unique to the Chinese populations, but some SNPs, such as rs13385731 (*RASGRP3*) and rs6590330 (*ETS1*), did show different allele frequencies between the two populations (**Supplementary Table 4b**).

The disparities of the GWAS findings between Chinese Han and European populations might suggest genetic heterogeneity of SLE between the two populations, especially in instances where risk alleles show a similar allele frequency in the two populations but association in only one population. Furthermore, although our study has $>90\%$ power to detect a SNP with MAF >0.2 and an OR >1.3 , it may not have sufficient power to detect risk SNPs that are less frequent and/or associated with weaker genetic effects, for example, several risk SNPs identified in European populations that were found to be very rare in Chinese populations. The genetic effects (ORs) of some associations may also be overestimated by previous studies in European populations. Consequently, some of the association disparities may simply be due to insufficient power and/or chance. In addition, some of the disparities could also be due to different LD and haplotype patterns between the two populations. It is possible that, although existing in both the Chinese Han and European populations, some of the causal variants are not sufficiently tagged by any SNPs screened in our GWAS analysis and thus failed to be detected in our study of the Chinese Han population. Furthermore, differences in clinical sub-phenotypes and/or environment risk factors between the two populations might also explain these disparities. Additional studies using large samples of diverse ethnic populations will help clarify the basis for disparities of SLE association between different populations and provide a better understanding of the genetic basis of SLE.

In summary, we performed a large GWAS of SLE in a Chinese Han population and identified nine new susceptibility loci, as well as confirming seven previously reported risk loci. Our study not only advances our understanding of the genetic basis of SLE susceptibility but also highlights the genetic heterogeneity of disease susceptibility between different ethnic populations.

METHODS

Methods and any associated references are available in the online version of the paper at <http://www.nature.com/naturegenetics/>.

Note: Supplementary information is available on the Nature Genetics website.

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AUTHOR CONTRIBUTIONS

X.-J.Z. conceived of this study and obtained financial support. X.-J.Z., J.-J.L. and S.Y. designed the study. J.-W.H., H.-F.Z., Y.C. and L.-D.S. participated in the design and were responsible for sample selection, genotyping and project management. D.-Q.Y., Z.H., Jin-Hua Xu, Z.-M.C., H.-F.X., H.F., Q.-J.L., Jian-Hua Xu, X.-P.L., Y.-F.P., D.-Q.D., F.-Q.Z., Z.-Z.Y., X.-Y.Z., Q.-W.W., F.H., L.M., X.-X.Z., W.-F.Z., F.G., P.-L.Z., Q.G., B.L., M.G., F.-L.X., J.-L.H., S.-X.L., H.L., Z.-X.W., C.-J.Y., P.-G.W., W.-M.Z., A.-P.Z., D.L. and Z.-F.Z. conducted sample selection and data management. W.-H.D., J.-Q.Y., S.-K.S., J.L., Y.-J.S., K.-J.Z., Yang Li, D.-Y.H. and Y.-M.L. undertook recruitment, collected phenotype data, undertook related data handling and calculation, managed recruitment and obtained biological samples. W.H., G.-P.Z., F.-S.Z., Y.-L.C., C.Q., C.Z., Z.Z., W.-S.L., Y.-Q.R., S.-Q.Z., M.S. and Y.W. performed genotyping analysis. H.-F.Z., X.-B.Z., Yi Li, H.Q.L. and F.-Y.Z. undertook data processing, statistical analysis and bioinformatics investigations. All the authors contributed to the final paper, with X.-J.Z., J.-J.L., S.Y., J.-W.H., H.-F.Z., Y.C. and L.-D.S. playing key roles.

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ONLINE METHODS

Subjects. All Chinese Han samples used in this study were obtained from doctors through collaboration with multiple hospitals in two geographic regions (central and southern China)^{22–24}. Samples in the initial stage (1,099 SLE cases and 1,254 controls) were recruited from central China; samples in the replication studies were recruited from central (replication 1: 1,643 cases and 5,930 controls) and southern (replication 2: 1,509 cases and 1,120 controls) China. All affected individuals defined as cases were diagnosed using the revised criteria for the classification of SLE from the American College of Rheumatology²⁵. Clinical information was collected from the affected individuals through a full clinical checkup by physician specialists. Additional demographic information was collected from both cases and controls through a structured questionnaire. All controls were clinically assessed to be without SLE, other autoimmune disorders, systemic disorders or family history of autoimmune disorders (including first-, second- and third-degree relatives). All participants provided written informed consent. The study was approved by the Institutional Ethical Committee of each hospital and was conducted according to Declaration of Helsinki principles.

EDTA anticoagulated venous blood samples were collected from all participants. Genomic DNA was extracted from peripheral blood lymphocytes by standard procedures using Flexi Gene DNA kits (Qiagen) and was diluted to working concentrations of 50 ng/μl for genome-wide genotyping and 15–20 ng/μl for the validation study.

Genotyping and quality controls in GWAS. The genome-wide genotyping analysis was conducted using Illumina Human610-Quad BeadChips at the Chinese National Human Genome Center at Shanghai, China. Genotyping was performed as previously described²⁶ according to the Infinium HD protocol from Illumina. Briefly, 200 ng of genomic DNA of each sample was whole-genome amplified, fragmented, precipitated and resuspended in the appropriate hybridization buffer. Denatured samples were hybridized on prepared Illumina Human610-Quad BeadChips. After hybridization, the BeadChip oligonucleotides were extended by a single labeled base, which was detected by fluorescence imaging with an Illumina Bead Array Reader. Normalized bead intensity data obtained for each sample were loaded into the Illumina BeadStudio 3.2 software, which converted the fluorescence intensities into SNP genotypes. The clustering of genotypes was carried out with the GenCall software version 6.2.0.4, which assigns a quality score to each locus and an individual genotype confidence score that is based on the distance of a genotype from the center of the nearest cluster.

All SNPs on the X, Y and mitochondrial chromosomes as well as copy number variation-related SNPs were excluded. SNPs were excluded if they showed either a call rate lower than 90% in cases or controls, a minor allele frequency <1% in the population or significant deviation from Hardy-Weinberg equilibrium in the controls ($P \leq 10^{-7}$). In the GWAS stage, 1,099 SLE cases and 1,254 controls were genotyped initially using the Illumina Human610-Quad BeadChip array. Fifty samples with genotyping call rate less than 98% were removed from analysis (33 cases and 17 controls). We then examined potential genetic relatedness based on pairwise identity by state for all the successfully genotyped samples using PLINK 1.06 software²⁷. In total, 51 samples were removed due to sample duplications and genetic relatedness (19 cases and 32 controls). After quality control, the genotype data of 493,955 autosomal SNPs in 1,047 cases and 1,205 controls were used for the GWAS analysis. The remaining samples were subsequently assessed for population stratification using the software package EIGENSTRAT²⁸. The original script from EIGENSTRAT was modified to extract the principal components for plotting (Supplementary Fig. 1).

SNP selection for replication. First, we chose SNPs with P value <0.01 after adjustment by gender and with P value < 5×10^{-4} without adjustment by gender for further evaluation. Then, all the chosen SNPs were further analyzed by including another 4,001 samples as controls. These samples were genotyped in a series of GWAS of various diseases in the Chinese Han population, including psoriasis, leprosy, vitiligo and atopic dermatitis. The improvement of association evidence by the inclusion of the 4,001 samples as controls was used as one

of the criteria for selecting SNPs for validation analysis; however, because these additional controls had diseases other than SLE, these 4,001 samples were not included in the association analysis of the validation study or in the final joint association analysis of the combined samples. Additional criteria for SNPs selection were high MAF (>0.05 both in cases and controls), high P value of Hardy-Weinberg equilibrium ($P \geq 0.01$ in controls) and proximity to putative candidate genes (immune-related or involved in immune cell proliferation and differentiation) or known susceptibility loci for SLE. In total, 78 SNPs were selected for the validation analysis, including 12 SNPs within or close to known susceptibility loci for SLE identified by previous studies (including *STAT4*, *TNFSF4*, *BLK*, *IRF5*, *TNFAIP3*, 6q21 and 22q11.21) (Table 2 and Supplementary Table 3).

Genotyping and quality controls in replication studies. Genotyping analysis of the SNPs selected for fast-track validation analysis was performed using the Sequenom MassArray system. Approximately 15 ng of genomic DNA was used to genotype each sample. Locus-specific PCR and detection primers were designed using the MassARRAY Assay Design 3.0 software (Sequenom) at the Key Laboratory of Dermatology at Anhui Medical University, Ministry of Education, China following the manufacturer's instructions. The DNA samples were amplified by multiplex PCR reactions, and the PCR products were then used for locus-specific single-base extension reactions. The resulting products were desalted and transferred to a 384-element SpectroCHIP array. Allele detection was performed using MALDI-TOF MS. The mass spectrograms were analyzed by the MassARRAY Typer software (Sequenom). To evaluate the quality of the genotype data for the validation analysis, 100 randomly selected samples from the GWAS stage were re-genotyped in replication samples by using the Sequenom system. The concordance rate between the genotypes from the Illumina and the Sequenom analyses was >99%. We excluded SNPs with call rate lower than 90% in cases or controls and with Hardy-Weinberg equilibrium in the controls ($P \leq 0.001$). After quality control, 9 SNPs with low call rates were excluded, and 69 SNPs were left for final analysis. The remaining SNP cluster patterns of the genotyping data from the Illumina and Sequenom analyses were checked to confirm their good quality.

Statistical analysis. In GWAS stage, single-marker association analyses were performed using logistic regression with gender as a covariate. For the replication studies, 69 SNPs that passed quality control were analyzed in each replication study using logistic regression with gender as a covariate. The joint analysis of all combined samples was performed using logistic regression with gender and sample cohorts as covariates. The P values adjusted by gender were reported without correction for multiple testing. Multiple logistic regression analysis was used to test the independence of multiple associations of top 13 SNPs within the MHC region. We used logistic regression analysis to test the independence of association evidence between MHC and non-MHC SNPs by controlling for the effect of rs9271100 and rs3997854 (adjusted for gender). To test for any two-locus interactions, we computed a logistic regression model for each pair of the top SNPs of the 16 non-MHC loci and 2 MHC SNPs (rs9271100 and rs3997854). Recombination plots of each newly discovered susceptibility locus were generated using the information from the HapMap project (CHB and JPT samples).

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